

def: Let $f_{ij} = P_i(\exists n \geq 1 : X_n = j) = \sum_{n=1}^{\infty} f_{ij}^{(n)}$.

Billingsley def $\left(= P_i\left(\bigcup_{n=1}^{\infty} \{X_n = j\}\right) \right)$ - Prob. of at least 1 visit

This is the probability, starting from i , that we ever hit state j .

def: A state $i \in S$ is recurrent (or persistent) if $f_{ii} = 1$, i.e. starting from i we are certain to return to i .

def: A state $i \in S$ is transient if it is not recurrent, i.e. if $f_{ii} < 1$.

Thm 8.2.1: Let $\{X_n : n \geq 0\}$ be a Markov chain on a state space S . Let $i \in S$. Then i is transient iff

$$P_i(X_n = i \text{ i.o.}) = 0 \text{ and iff } \sum_{n=1}^{\infty} P_{ii}^{(n)} < \infty.$$

On the other hand, i is recurrent iff

$$P_i(X_n = i \text{ i.o.}) = 1 \text{ and iff } \sum_{n=1}^{\infty} P_{ii}^{(n)} = \infty.$$

def: $\left[P_{ij}^{(n)}$ is the n^{th} step transition probability: $P(X_{m+n} = j \mid X_m = i)$]

To prove this Thm, we first need a lemma:

Lemma 8.2.2 : $P_i(\#\{n \geq 1 : X_n = j\} \geq k) = f_{ij} (f_{jj})^{k-1}$,
 for $k=1,2,\dots$ In words, the probability that, starting from i , we hit state j at least k times, is $f_{ij} (f_{jj})^{k-1}$.
~~It seems to be a typo~~

Pf: Starting from i , hitting j at least k times is equiv. to first hitting j (from i) & then returning to j at least $k-1$ more times. [See B p. 117 for more details]

Pf of Thm 8.2.1: By continuity of prob. & by Lemma 8.2.2,

$$P_i(X_n = i \text{ i.o.}) = \lim_{k \rightarrow \infty} P_i(\#\{n \geq 1 : X_n = i\} \geq k)$$

$$= \lim_{k \rightarrow \infty} (f_{ii})^k = \begin{cases} 0 & \text{if } f_{ii} < 1 \\ 1 & \text{if } f_{ii} = 1 \end{cases}$$

This proves the first equivalence to transience & recurrence.

For the remaining 2 equivalences,

$$\sum_{n=1}^{\infty} P_{ii}^{(n)} = \sum_{n=1}^{\infty} P_i(X_n = i) \quad \text{by def}$$

$$= \sum_{n=1}^{\infty} E_i(\mathbb{1}_{X_n = i})$$

$$\begin{aligned}
 &= E_i \left(\sum_{n=1}^{\infty} \mathbb{1}_{X_n=i} \right) \quad \text{by countable linearity} \\
 &= E_i \left(\#\{n \geq 1 : X_n=i\} \right) \\
 &= \sum_{k=1}^{\infty} P_i \left(\#\{n \geq 1 : X_n=i\} \geq k \right) \quad \text{by Prop. 4.2.9} \\
 &= \sum_{k=1}^{\infty} (f_{ii})^k \quad \text{by Lemma 0.2.2} \\
 &= \begin{cases} \frac{f_{ii}}{1-f_{ii}} < \infty & \text{if } f_{ii} < 1 \quad (\text{geometric series}) \\ \infty & \text{if } f_{ii} = 1 \quad \square \end{cases}
 \end{aligned}$$

$\left(\sum_{k=1}^{\infty} P(X \geq k) = E[LX] \right)$
 for X nonneg RV

• Apply this Thm to ^{simple} symmetric RW ($p = \frac{1}{2}$):

$S = \mathbb{Z}, X_0 = 0$

For any $i \in \mathbb{Z}$, $P_{ii}^{(n)} = \binom{n}{n/2} \left(\frac{1}{2}\right)^n$ for n even

(and $P_{ii}^{(n)} = 0$ for n odd). Sterling's approx for $n!$

$\Rightarrow P_{ii}^{(n)} \sim \sqrt{2/\pi n}$ for large n

Thus, $\sum_{n=1}^{\infty} P_{ii}^{(n)} = \infty$ and all states are recurrent.

$\left(\sqrt{n} \leq n \text{ for } n \geq 1 \Rightarrow \frac{1}{\sqrt{n}} \geq \frac{1}{n} \text{ \& } \sum \frac{1}{n} \text{ diverges so } \sum \frac{1}{\sqrt{n}} \text{ diverges as well} \right)$

half steps R
half steps L

- Apply to asymmetric RW (simple) ($p \neq \frac{1}{2}$):

$$S = \mathbb{Z}$$

$$P_{ii}^{(n)} = \binom{n}{n/2} p^{n/2} (1-p)^{n/2} \text{ with } p(1-p) < \frac{1}{4}.$$

$$\text{sterling} \Rightarrow P_{ii}^{(n)} \sim [4p(1-p)]^{n/2} \sqrt{2/\pi n} \text{ with } 4p(1-p) < 1$$

Thus, $\sum_{n=1}^{\infty} P_{ii}^{(n)} < \infty$ and all states are transient.

[You can fill in the details here!]

$$\left[\sqrt{\frac{2}{\pi}} \sum_{n=1}^{\infty} \frac{(4p(1-p))^{n/2}}{\sqrt{n}} < \infty \text{ by the ratio test} \right. \\ \left. \text{(see next pg. scratch notes)} \right]$$

In higher dimensions, $S = \mathbb{Z}^d \leftarrow d\text{-dim lattice}$

symmetric RW is recurrent in dimensions $1 \leq 2$,
but transient in dimensions $d \geq 3$.

↗
counter-intuitive!

$$\text{details} \left[\begin{array}{l} P_{ii}^{(n)} \sim \left(\sqrt{2/\pi n}\right)^d = (2/\pi n)^{d/2} \\ \sum_{n=1}^{\infty} P_{ii}^{(n)} = \infty \text{ iff } d \leq 2 \rightarrow p\text{-test} \end{array} \right]$$

p-series test

$$\sum_{n=1}^{\infty} \frac{1}{n^p} \begin{cases} \text{converges if } p > 1 \\ \text{diverges if } p \leq 1 \end{cases}$$

$$\sum_{n=1}^{\infty} \underbrace{[4p(1-p)]^{n/2}}_{a < 1} \sqrt{\frac{2}{\pi n}} < \infty$$

Q. Why does this converge?

$$\sqrt{\frac{2}{\pi}} \sum a^{n/2} \cdot \frac{1}{n^{1/2}}$$

$$a_n = \left(\frac{a^n}{n}\right)^{1/2}$$

Ratio Test *

$$\lim_{n \rightarrow \infty} \left| \frac{a_{n+1}}{a_n} \right| = L < 1 \Rightarrow \sum_{n=1}^{\infty} a_n \text{ absol. convergent}$$

$$\lim_{n \rightarrow \infty} \frac{\left(\frac{a^{n+1}}{n+1}\right)^{1/2}}{\left(\frac{a^n}{n}\right)^{1/2}} = \lim_{n \rightarrow \infty} \left(\frac{a^n a}{n+1} \cdot \frac{n}{a^n}\right)^{1/2} = \lim_{n \rightarrow \infty} a^{1/2} \underbrace{\left(\frac{n}{n+1}\right)^{1/2}}_{\frac{\sqrt{n}}{\sqrt{n+1}}}$$

L'Hopital's Rule \rightarrow $\stackrel{H}{=} a^{1/2} \lim_{n \rightarrow \infty} \frac{\sqrt{n+1} \cdot \frac{1}{2} n^{-1/2} - \sqrt{n} \cdot \frac{1}{2} (n+1)^{-1/2}}{n+1}$

To use `cutree` with `agnes` and `diana` you can perform the following:

```
# Cut agnes() tree into 4 groups
hc_a <- agnes(df, method = "ward")
cutree(as.hclust(hc_a), k = 4)
```

```
# Cut diana() tree into 4 groups
hc_d <- diana(df)
cutree(as.hclust(hc_d), k = 4)
```

$$\frac{\frac{1}{2} \sqrt{n+1} \cdot \frac{1}{\sqrt{n}} - \frac{1}{2} \sqrt{n} \cdot \frac{1}{\sqrt{n+1}}}{n+1}$$

$$= \frac{\frac{1}{2} \left(\frac{\sqrt{n+1}}{\sqrt{n}} - \frac{\sqrt{n}}{\sqrt{n+1}} \right)}{n+1}$$

$$= \frac{1}{2} \left(\frac{(\sqrt{n+1})^2 - (\sqrt{n})^2}{\sqrt{n} \sqrt{n+1} (n+1)} \right)$$

Lastly, we can also compare two dendrograms. Here we compare hierarchical clustering with complete linkage versus Ward's method. The function `tanglegram` plots two dendrograms, side by side, with their labels connected by lines = $\frac{1}{2} (n+1 - n)$

```
# Compute distance matrix
res.dist <- dist(df, method = "euclidean")
```

```
# Compute 2 hierarchical clusterings
hc1 <- hclust(res.dist, method = "complete")
hc2 <- hclust(res.dist, method = "ward.D2")
```

```
# Create two dendrograms
dend1 <- as.dendrogram(hc1)
dend2 <- as.dendrogram(hc2)
```

```
tanglegram(dend1, dend2)
```

$$= \frac{\frac{1}{2} (n+1 - n)}{\sqrt{n(n+1)} (n+1)} = \frac{1}{2\sqrt{n(n+1)} (n+1)}$$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{1}{2\sqrt{n(n+1)}^{3/2}} = 0$$

$$\Rightarrow \sum_{n=1}^{\infty} a_n \text{ converges!}$$



def: A Markov chain is irreducible if $f_{ij} > 0$

$\forall i, j \in S$, i.e. if it is possible to move from any state to any other state.

Equivalently, MC is irreducible if for any $i, j \in S$,
 $\exists n \in \mathbb{N}$ st. $P_{ij}^{(n)} > 0$.

def: A MC is reducible if it is not irreducible, i.e.
 if $f_{ij} = 0$ for some $i, j \in S$.

Thm 8.2.3: Let $P = \{P_{ij}\}_{i, j \in S}$ be the transition probability matrix for an irreducible Markov chain on a state space S . Then the following are equivalent:

- (i) $\exists k \in S$ with $f_{kk} = 1$, i.e. k is recurrent.
- (ii) $\forall i, j \in S$, we have $f_{ij} = 1 \rightarrow$ All states are recurrent.
- (iii) $\exists k, l \in S$ with $\sum_{n=1}^{\infty} P_{kl}^{(n)} = \infty$.
- (iv) $\forall i, j \in S$, we have $\sum_{n=1}^{\infty} P_{ij}^{(n)} = \infty$.

Pf: Clearly (ii) \Rightarrow (i). We see that (i) \Rightarrow (iii) follows immediately from Thm 8.2.1.

$2 \Rightarrow 1 \Rightarrow 3 \Rightarrow 4$

Next show that (iii) \Rightarrow (iv):

Assume (iii) holds. Let $i, j \in S$. By irreducibility,

$\exists m, r \in \mathbb{N}$ s.t. $P_{ik}^{(m)} > 0$ and $P_{lj}^{(r)} > 0$. But then,

since $P_{ij}^{(m+n+r)} \geq P_{ik}^{(m)} P_{kl}^{(n)} P_{lj}^{(r)}$, it follows that

$$\sum_n P_{ij}^{(n)} \geq P_{ik}^{(m)} P_{lj}^{(r)} \underbrace{\sum_n P_{kl}^{(n)}}_{\text{from (iii)}} = \infty \text{ as claimed.}$$

Lastly, show that (iv) \Rightarrow (ii):

[see next pg]