

Before we prove Thm 9.3.4, need a lemma.

Lemma 9.3.5 : Let Z be a RV with $E[Z] < 0$ s.t.

$M_Z(s) < \infty$ for $-a < s < b$, for some $a, b > 0$. Then

$$P(Z \geq 0) \leq \rho \text{ where } \rho = \inf_{0 < s < b} M_Z(s) < 1.$$

Pf: For any $s \in (0, b)$ the function $f(x) = e^{sx}$ is increasing, so by Markov's Ineq.

$$P(Z \geq 0) = P(e^{sZ} \geq 1) \leq \frac{E[e^{sZ}]}{1} = M_Z(s)$$

Now take infimum over $0 < s < b$:

$$P(Z \geq 0) \leq \inf_{0 < s < b} M_Z(s) = \rho.$$

Lastly, since $M_Z(0) = 1$ and $M'_Z(0) = E[Z] < 0$ (by assumption), it follows that $M_Z(s) < 1$ for all positive s sufficiently close to 0, i.e. $\rho < 1$. (calculus)

c21 la Pf of Thm 9.3.4 : Let $Y_i = X_i - \mu - \varepsilon$ s.t. $E[Y_i] = -\varepsilon < 0$.

For $-a < s < b$, we have that

$$M_{Y_i}(s) = E[e^{sY_i}] = e^{-s(\mu+\varepsilon)} E[e^{sX_i}] = e^{-s(\mu+\varepsilon)} M_{X_i}(s) < \infty$$

(since $M_{X_i}(s) < \infty \forall i$ by assumption).

Then by Lemma 9.3.5,

$$P\left(\frac{X_1 + \dots + X_n}{n} > \mu + \varepsilon\right) = P\left(\frac{Y_1 + \dots + Y_n}{n} \geq 0\right)$$

$$= P(Y_1 + \dots + Y_n \geq 0)$$

$$\leq \inf_{0 < s < b} M_{Y_1 + \dots + Y_n}(s) = \inf_{0 < s < b} (M_{Y_1}(s))^n = p^n$$

↑
since Y_i 's
are indep

$$\text{where } p = \inf_{0 < s < b} M_{Y_1}(s) = \inf_{0 < s < b} (e^{-s(\mu+\varepsilon)} M_{X_1}(s)).$$

Lastly, $p < 1$ by Lemma 9.3.5. ◻

Weak Convergence

(Ref: R § 10, B § 25)

def: Given Borel probability distributions μ, μ_1, μ_2, \dots on \mathbb{R} , $\{\mu_n\}$ converges weakly to μ if

$$\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu \text{ for all bounded continuous functions } f: \mathbb{R} \rightarrow \mathbb{R}.$$

Denoted by $\boxed{\mu_n \Rightarrow \mu}$.

Recall: The boundary of a set $A \subseteq \mathbb{R}$ is

$$\partial A = \{x \in \mathbb{R} : \forall \varepsilon > 0, A \cap (x-\varepsilon, x+\varepsilon) \neq \emptyset, A^c \cap (x-\varepsilon, x+\varepsilon) \neq \emptyset\}.$$

Thm: The following are equivalent:

(1) $\mu_n \Rightarrow \mu$ (weak convergence)(2) $\mu_n(A) \rightarrow \mu(A)$ \forall measurable sets A s.t. $\mu(\partial A) = 0$.(3) $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$ $\forall x \in \mathbb{R}$ s.t. $\mu(\{x\}) = 0$.(4) [Skorohod's Thm]: \exists RVs Y, Y_1, Y_2, \dots defined jointly on some (Ω, \mathcal{F}, P) with $L(Y) = \mu$ and $L(Y_n) = \mu_n \quad \forall n \in \mathbb{N}$ s.t. $Y_n \rightarrow Y$ with prob. 1.

(5) $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$ & bounded Borel-measurable functions

$f: \mathbb{R} \rightarrow \mathbb{R}$ s.t. $\mu(D_f) = 0$ where D_f is the set of points where f is discontinuous.

PF: Exercise to the reader! (Rosenthal)

Example 1: Let μ be Lebesgue measure on $[0,1]$

let $\mu_n(\frac{i}{n}) = \frac{1}{n}$ for $i=1, 2, \dots, n$.

μ = continuous

μ_n = discrete

Note that $\mu(Q) = 0$ while $\mu_n(Q) = 1$ for each n .

We do have $\mu_n \rightarrow \mu$ as $n \rightarrow \infty$. Why?

For any $0 \leq x \leq 1$, $\mu((-\infty, x]) = x$

while $\mu_n((-\infty, x]) = \frac{\lfloor nx \rfloor}{n}$.

$$\text{Then } |\mu_n((-\infty, x]) - \mu((-\infty, x])| = \left| \frac{\lfloor nx \rfloor}{n} - x \right|$$

$$\leq \frac{1}{n} \rightarrow 0 \text{ as } n \rightarrow \infty$$

Note that $\sigma(Q \cap [0,1]) = [0,1]$ so $\mu(\sigma(Q \cap [0,1])) \neq 0$
 so doesn't contradict part (2).

Connections to other types of convergence

Prop: If $X_n \rightarrow X$ in probability, then $\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X)$ as $n \rightarrow \infty$.

Remark: Sometimes write $\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X)$ as $n \rightarrow \infty$ as $X_n \rightarrow X$ as $n \rightarrow \infty$. We say that $\{X_n\}$ converges weakly to X , or in distribution.

$$X_n \xrightarrow{\text{a.s.}} X \implies X_n \xrightarrow{P} X \implies X_n \Rightarrow X \text{ (in dist'n)} \\ (\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X))$$

$\Rightarrow \exists$ R.V.s $Y_n \in Y$ having the same laws
s.t. $Y_n \xrightarrow{\text{a.s.}} Y$ (Thm, (4))

Note: The converse of the proposition above is FALSE.

$$\mathcal{L}(X_n) = \mathcal{L}(X) \not\Rightarrow X_n \xrightarrow{P} X \text{ as } n \rightarrow \infty,$$

Example: X, X_1, X_2, \dots i.i.d. each equal to ± 1 with prob. $\frac{1}{2}$, then $\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X)$ but

$$P(|X - X_n| \geq 2) = \frac{1}{2} \not\rightarrow 0 \text{ as } n \rightarrow \infty$$

either $X=1$ and $X_n=-1 \rightarrow$ prob $\frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$ } sum these
or $X=-1$ and $X_n=+1 \rightarrow$ prob $\frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$ } to get $\frac{1}{2}$

so X_n does NOT converge to X in probability or w/prob. 1.

[However, if X is constant then the converse does hold.]

Lastly, we can use Skorohod's Thm to translate results involving convergence w/prob. 1 to results involving weak convergence (or convergence in prob. by using prop 10.2.1 R.)

For example,

prop: Suppose $\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X)$ as $n \rightarrow \infty$, with $X_n \geq 0$.

Then $E[X] \leq \liminf_n E[X_n]$.

[pf uses Fatou's Lemma].

e.g. If $X = 0$ and $P(X_n = n) = \frac{1}{n}$, $P(X_n = 0) = 1 - \frac{1}{n}$,

then $\mathcal{L}(X_n) \Rightarrow \mathcal{L}(X)$ and

$$0 = E[X] \leq \liminf_n E[X_n] = 1$$

Characteristic Functions

(Ref: R § 11, B § 26) — this will not be on Exam 2,
yes on the Final

def: Given a random variable X , its characteristic function (aka Fourier Transform) is given by

$$\boxed{\phi_X(t) = E[e^{itX}]} = E[\cos(tx)] + iE[\sin(tx)], t \in \mathbb{R}$$

↑

- Function from \mathbb{R} to \mathbb{C}
- By change of variable thm, $\phi_X(t)$ only depends on the distribution of X

"Similar" to MGF ($M_X(s) = E[e^{sx}]$) but the appearance of the imaginary number $i = \sqrt{-1}$ is a significant difference:

$$|e^{itX}| = 1 \text{ for any real } t \text{ and } X$$

$$\begin{aligned} e^{it} &= \cos t + i \sin t \\ |e^{it}| &= \sqrt{\cos^2 t + \sin^2 t} \\ &= 1 \end{aligned} \quad \text{for } t \in \mathbb{R}$$

$$|\phi_X(t)| \leq 1 < \infty \quad \forall t, X \quad \text{whereas}$$

Always exists
since it is
bounded

$M_X(s)$ can be infinite
for some $s \neq 0$.

More details :

A RV X with distribution μ has characteristic function

$$\phi_X(t) = E[e^{itX}] = \int_{-\infty}^{\infty} e^{itx} d\mu(x) = \int_{-\infty}^{\infty} e^{itx} f_X(x) dx$$

↗ If X has PDF f_X
that is Riemann integrable

Properties

- The characteristic function uniquely determines the distribution (as in MGFs).
- Pointwise convergence of characteristic functions implies weak convergence of the corresponding distributions.
→ Continuity Thm,
used to prove CLT

Moments & Derivatives

As with MGFs,

- $\phi_X(0) = 1$ for any RV X
- $X \neq Y$ indep RVs $\Rightarrow \phi_{X+Y}(t) = \phi_X(t) \phi_Y(t)$

Let $\mu = \mathcal{L}(X)$. Then

$$|\phi_X(t+h) - \phi_X(t)| = \left| \int (e^{i(t+h)x} - e^{itx}) d\mu(x) \right|$$

Lecture 6

1 Characteristic Functions

For weak convergence of probability measures on \mathbb{R}^d , or equivalently weak convergence of \mathbb{R}^d -valued random variables, an important *convergence determining class* of functions are functions of the form $f(x) = e^{it \cdot x}$, with $t \in \mathbb{R}^d$. They lead to

Definition 1.1 [Characteristic Function] Let X be a \mathbb{R}^d -valued random variable with distribution μ on \mathbb{R}^d . Then the characteristic function of X (or μ) is defined to be

$$\phi(t) := \mathbb{E}[e^{it \cdot X}] = \int e^{it \cdot x} \mu(dx).$$

When μ has a density with respect to Lebesgue measure, i.e., $\mu(dx) = \rho(x)dx$, $\phi(\cdot)$ is just the Fourier transform of the function $\rho(\cdot)$. Therefore in general, we can think of $\phi(\cdot)$ as the Fourier transform of the measure μ .

Here are some properties which are immediate from the definition:

Proposition 1.2 [Properties of Characteristic Functions]

- (i) Since $e^{it \cdot x} = \cos tx + i \sin tx$ has bounded real and imaginary parts, $\phi(t)$ is well-defined.
 - (ii) $\phi(0) = 1$ and $|\phi(t)| \leq 1$ for all $t \in \mathbb{R}^d$.
 - (iii) ϕ is uniformly continuous on \mathbb{R}^d . More precisely,
- $$|\phi(t+h) - \phi(t)| = |\mathbb{E}[e^{i(t+h) \cdot X}] - \mathbb{E}[e^{it \cdot X}]| = |\mathbb{E}[e^{it \cdot X}(e^{ih \cdot X} - 1)]| \leq \mathbb{E}[|e^{ih \cdot X} - 1|],$$
- which tends to 0 as $h \rightarrow 0$ by the bounded convergence theorem.
- (iv) The complex conjugate of ϕ , $\overline{\phi(t)}$, is the characteristic function of $(-X)$, and $\phi(t) \in \mathbb{R}$ for all $t \in \mathbb{R}^d$ if $X \stackrel{\text{dist}}{=} (-X)$.
 - (v) For $a \in \mathbb{R}$ and $b \in \mathbb{R}^d$, $aX + b$ has characteristic function $\varphi(t) = e^{ib \cdot t} \phi(at)$.
 - (vi) If X and Y are two independent random variables with characteristic functions ϕ_X and ϕ_Y , then $X + Y$ has characteristic function $\phi_{X+Y}(t) = \mathbb{E}[e^{it \cdot (X+Y)}] = \phi_X(t)\phi_Y(t)$.

Proposition 1.3 [Common Distributions and Their Characteristic Functions]

Examples

- (a) The delta measure at $a \in \mathbb{R}$: $\mu(dx) = \delta_a(dx)$. Then $\phi(t) = e^{ia \cdot t}$.
- (b) The coin flip: $\mu(\{1\}) = \mu(\{-1\}) = \frac{1}{2}$. Then $\phi(t) = \frac{e^{it} + e^{-it}}{2} = \cos t$.
- (c) The Poisson distribution with parameter $\lambda > 0$: $\mu(\{n\}) = e^{-\lambda} \frac{\lambda^n}{n!}$ for $n \in \{0\} \cup \mathbb{N}$. Then $\phi(t) = e^{-\lambda} \sum_{n=0}^{\infty} e^{itn} \frac{\lambda^n}{n!} = e^{-\lambda(1-e^{it})}$.
- (d) The compound Poisson distribution: $X = \sum_{i=1}^T Y_i$, where T has Poisson distribution with parameter λ , and $(Y_i)_{i \in \mathbb{N}}$ is an independent sequence of i.i.d. random variables with characteristic function $\psi(\cdot)$. Then $\phi(t) = \mathbb{E}[e^{itX}] = e^{-\lambda(1-\psi(t))}$.

$$\leq \int |e^{i(t+h)x} - e^{itx}| d\mu(x)$$

$$|ab| = |a||b|$$

$$= \underbrace{\int |e^{itx}| |e^{ihx} - 1| d\mu(x)}_1 = \underbrace{\int |e^{ihx} - 1| d\mu(x)}$$

As $h \rightarrow \infty$, this decreases to 0 by the Bounded Convergence Thm (since $|e^{ihx} - 1| \leq 2$). Hence, $\phi_x(t)$ is always a uniformly continuous function.

Derivatives of $\phi_x(t)$ are straightforward.

Prop 11.0.1: Suppose X is a RV with $E[|X|^k] < \infty$. Then for $0 \leq j \leq k$, ϕ_x has finite j^{th} derivative given by

$$\phi_x^{(j)}(t) = E[(ix)^j e^{itX}].$$

In particular, evaluate at $t=0$ to get

$$\boxed{\phi_x^{(j)}(0) = i^j E[X^j].}$$

PF: Uses induction on j . See text for details.